

Modified Ratio Estimator in Simple Random Sampling using Auxiliary Information

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ABSTRACT

In this study, we offer a new estimator for population variance in simple random sampling based on auxiliary information. We calculated the proposed estimator's bias and MSE equations and compared them to the bias and MSE of existing estimators, demonstrating that the new estimator is more efficient than the existing estimators proposed by different authors. With the aid of numerical example, we can support this theoretical result.

Keywords

Variance estimator, bias, MSE, simple random sampling, auxiliary information, efficiency

1. INTRODUCTION

Auxiliary information has been used extensively in estimation of parameters like mean and variance since several decades. Isaki (1983) [1] got inspiration from ratio estimator of finite population mean and proposed a ratio estimator, usually called classical estimator, of finite population variance. But Singh and Solanki (2013a) claimed that Isaki (1983) [1] ratio estimator is the member of the class of estimators developed by Das and Tripathi (1978) [3]. Arcos and Rueda (1997) [4] suggested multivariate ratio estimator for population variance. Ahmed et al. (2000) [5] criticized the claim of Arcos and Rueda (1997) [4], Kadilar and Cingi (2006a) [6] developed an estimator, ratio-type estimator of the mean of population, Kadilar and Cingi (2006b) [7] extended the idea of Isaki (1983) [1] ratio estimator for population variance. Kadilar and Cingi (2006b) [7] involved the information available about coefficient of variation and coefficient of kurtosis of the auxiliary variable to generate these estimators under simple random sampling as well. Gupta and Shabbir (2008) [8] gave a hybrid class of variance estimators of population mean. Subramani and Kumarapandiyan (2012a) [9] modified the usual ratio-type estimator of Kadilar and Cingi (2006b) [7] for population variance using population median obtained from auxiliary variable. Subramani and Kumarapandiyan (2012b) [10] further modified the usual ratio-type variance estimators using lower and upper quartiles, inter-quartile range, quartile deviation and quartile average of the auxiliary variable. Subramani and Kumarapandiyan (2013) [11] developed another more efficient modified ratio-type estimator using median and coefficient of variation of the auxiliary variable.

2. NOTATION

Before delving deeper into the classical ratio type variance estimator, modified ratio type variance estimators, and proposed modified ratio type variance estimator, the notations that will be used in this paper are listed below:

1. Population size

N

2. Sample size

3. Population mean of the study variable y

$$\bar{Y} = \frac{\sum_{i=1}^N y_i}{N}$$

4. Population mean of the auxiliary variable x

$$\bar{X} = \frac{\sum_{i=1}^N x_i}{N}$$

5. Sample mean of the study variable y

$$\bar{y} = \frac{\sum_{i=1}^n y_i}{n}$$

6. Sample mean of the auxiliary variable x

$$\bar{x} = \frac{\sum_{i=1}^n x_i}{n}$$

7. Population variance of the study variable y

$$S_Y^2 = \frac{\sum_{i=1}^N (y_i - \bar{Y})^2}{N - 1}$$

8. Population variance of the auxiliary variable x

$$S_X^2 = \frac{\sum_{i=1}^N (x_i - \bar{X})^2}{N - 1}$$

9. Sample variance of the study variable y

$$S_y^2 = \frac{\sum_{i=1}^n (y_i - \bar{y})^2}{n - 1}$$

10. Sample variance of the auxiliary variable x

$$S_x^2 = \frac{\sum_{i=1}^n (x_i - \bar{x})^2}{n - 1}$$

11. Coefficient of variation

$$C_x = \frac{S_x}{\bar{X}}, \quad C_y = \frac{S_y}{\bar{Y}}$$

12. Skewness of the auxiliary variable

$$\beta_{1(x)} = \frac{\mu_{03}^2}{\mu_{02}^3}$$

13. Kurtosis of the auxiliary variable

$$\beta_{2(x)} = \frac{\mu_{04}^2}{\mu_{02}^2}$$

14. Kurtosis of the study variable

$$\beta_{2(y)} = \frac{\mu_{40}^2}{\mu_{20}^2}$$

15. $\gamma = \frac{(1-f)}{n}$

3. SOME EXISTING ESTIMATORS

1. Isaki (1983)

Motivated by the estimator of the population mean

$$\hat{\mu}_{yR} = \frac{\bar{y}}{\bar{x}} \bar{X}$$

Isaki (1983) [1] suggested the ratio estimators of the variance population using auxiliary information.

$$t_{Isaki} = s_y^2 \frac{S_x^2}{S_x^2}$$

The MSE of the above estimator using first order approximation is

$$MSE(t_{Isaki}) \cong \lambda' S_y^4 \{\beta_2(y) + \beta_2(x) - 2h\}$$

Or

$$MSE(t_{Isaki}) \cong \lambda' S_y^4 \{\beta_2'(y) + \beta_2'(x) - 2h'\}$$

And the bias is

$$Bias(t_{Isaki}) \cong \lambda' S_y^2 \{\beta_2(x) - h\}$$

Or

$$Bias(t_{Isaki}) \cong \lambda' S_y^2 \{\beta_2'(x) - h'\}$$

Or

$$Bias(t_{Isaki}) \cong -Cov\left(\frac{S_y^2}{S_x^2}, S_x^2\right)$$

2. Kadilar and Cingi (2005)

Kadilar and Cingi(2005) [2] proposed the following class of estimators for population variance:

$$t_{CH1} = s_y^2 [s_x^2 - \beta_2(x)]^{-1} [S_x^2 - \beta_2(x)]$$

$$t_{CH2} = s_y^2 [s_x^2 C_x - \beta_2(x)]^{-1} [S_x^2 C_x - \beta_2(x)]$$

$$t_{CH3} = s_y^2 (s_x^2 - C_x)^{-1} (S_x^2 - C_x)$$

$$t_{CH4} = s_y^2 [s_x^2 \beta_2(x) - C_x]^{-1} [S_x^2 \beta_2(x) - C_x]$$

The MSE's of the above estimators are given below respectively:

$$MSE(t_{CH1}) \cong \lambda' S_y^4 \left[\beta_2'(y) - 2 \frac{S_x^2}{S_x^2 - \beta_2(x)} h' + \left\{ \frac{S_x^2}{S_x^2 - \beta_2(x)} \right\}^2 \beta_2'(x) \right]$$

$$MSE(t_{CH2}) \cong \lambda' S_y^4 \left[\beta_2'(y) - 2 \frac{S_x^2}{S_x^2 C_x - \beta_2(x)} h' + \left\{ \frac{S_x^2}{S_x^2 C_x - \beta_2(x)} \right\}^2 \beta_2'(x) \right]$$

$$MSE(t_{CH3}) \cong \lambda' S_y^4 \left[\beta_2'(y) - 2 \frac{S_x^2}{S_x^2 - C_x} h' + \left\{ \frac{S_x^2}{S_x^2 - C_x} \right\}^2 \beta_2'(x) \right]$$

$$MSE(t_{CH4}) \cong \lambda' S_y^4 \left[\beta_2'(y) - 2 \frac{S_x^2}{S_x^2 \beta_2(x) - C_x} h' + \left\{ \frac{S_x^2}{S_x^2 \beta_2(x) - C_x} \right\}^2 \beta_2'(x) \right]$$

The Biases of above estimators are given below respectively:

$$Bias(t_{CH1}) \cong \lambda' S_y^2 \frac{S_x^2}{S_x^2 - \beta_2(x)} \left[\frac{S_x^2}{S_x^2 - \beta_2(x)} \beta_2'(x) - h' \right]$$

$$Bias(t_{CH2}) \cong \lambda' S_y^2 \frac{S_x^2}{S_x^2 C_x - \beta_2(x)} \left[\frac{S_x^2}{S_x^2 C_x - \beta_2(x)} \beta_2'(x) - h' \right]$$

$$Bias(t_{CH3}) \cong \lambda' S_y^2 \frac{S_x^2}{S_x^2 - C_x} \left[\frac{S_x^2}{S_x^2 - C_x} \beta_2'(x) - h' \right]$$

$$Bias(t_{CH4}) \cong \lambda' S_y^2 \frac{S_x^2}{S_x^2 \beta_2(x) - C_x} \left[\frac{S_x^2}{S_x^2 \beta_2(x) - C_x} \beta_2'(x) - h' \right]$$

4. PROPOSED ESTIMATOR

Using the known value of the auxiliary variable's population mean, we introduced a new modified ratio type variance estimator. For population variance, the modified ratio type variance estimator is defined as follows:

$$t_{new} = s_y^2 \frac{\rho^2 S_x^2 + \bar{X}^2}{\rho^2 s_x^2 + \bar{X}^2}$$

Bias of the proposed estimator

$$t_{new} = S_y^2 (1 + e_0) \frac{\rho^2 S_x^2 + \bar{X}^2}{\rho^2 S_x^2 (1 + e_1) + \bar{X}^2}$$

$$t_{new} = S_y^2 (1 + e_0) \{1 + \Omega e_1\}^{-1}$$

Where

$$\Omega = \frac{\rho^2 S_x^2}{\rho^2 S_x^2 + \bar{X}^2}$$

Expanding the expression by using Taylor's series

$$t_{new} = S_y^2 (1 + e_0) \{1 - \Omega e_1 + (\Omega e_1)^2 - (\Omega e_1)^3 + \dots\}$$

$$t_{new} = S_y^2 (1 + e_0) \{1 - \Omega e_1 + (\Omega e_1)^2\}$$

$$t_{new} - S_y^2 = S_y^2 \{e_0 - \Omega e_1 + (\Omega e_1)^2 - \Omega e_0 e_1\}$$

$$E(t_{new} - S_y^2) = E[S_y^2 \{e_0 - \Omega e_1 + (\Omega e_1)^2 - \Omega e_0 e_1\}]$$

$$B(t_{new}) = \Omega \lambda' S_y^2 \{\Omega \beta_2'(x) - h'\}$$

MSE of the proposed estimator

Now comes to MSE of the estimator

$$\begin{aligned}
 MSE(t_{new}) &= \{S_y^2(e_0 - \Omega e_1)\}^2 \\
 (\square \text{ ignore higher order}) \\
 \\
 MSE(t_{new}) &= S_y^4 E\{(e_0^2) + \Omega^2(e_1^2) - 2\Omega(e_0 e_1)\} \\
 \\
 &= S_y^4 \{\lambda' \beta'_{2(y)} + \Omega^2 \lambda' \beta'_{2(x)} - 2\Omega \lambda' h'\} \\
 \\
 &= \lambda' S_y^4 \left\{ \beta'_{2(y)} + \beta'_{2(x)} \left(\Omega^2 - 2\Omega \frac{h'}{\beta'_{2(x)}} \right) \right\} \\
 &= \lambda' S_y^4 \{\beta'_{2(y)} + \beta'_{2(x)}(\Omega^2 - 2\Omega v')\} \rightarrow (i)
 \end{aligned}$$

Where

$$\square \quad v' = \frac{h'}{\beta'_{2(x)}}$$

In order to minimize MSE, Differentiate (i) partially w.r.t Ω and equating to zero.

$$\frac{\partial}{\partial \Omega} MSE(t_{new}) = 0$$

$$\lambda' S_y^4 \frac{\partial}{\partial \Omega} \{\beta'_{2(y)} + \beta'_{2(x)}(\Omega^2 - 2\Omega v')\} = 0$$

$$\Omega = v'$$

So,

$$\begin{aligned}
 &= \lambda' S_y^4 \{\beta'_{2(y)} - \beta'_{2(x)} v'^2\} \\
 \\
 &= \lambda' S_y^4 \beta'_{2(y)} \left\{ 1 - v'^2 \frac{\beta'_{2(x)}}{\beta'_{2(y)}} \right\} \\
 \\
 &= \lambda' S_y^4 \beta'_{2(y)} \left\{ 1 - \left(\frac{h'}{\beta'_{2(x)} \beta'_{2(y)}} \right)^2 \right\} MSE(t_{new}) \\
 &= \lambda' S_y^4 \beta'_{2(y)} \{1 - (\rho')^2\}
 \end{aligned}$$

Where

$$\rho' = \frac{h'}{\sqrt{\beta'_{2(x)} \beta'_{2(y)}}}$$

5. NUMERICAL STUDY

The execution of the proposed modified ratio type variance estimator is surveyed with that of traditional ratio type estimator and existing modified ratio type variance estimators recorded in table 1 for certain common populations. The population parameters and the constants computed are given below:

Table 1: Parameters and Constants of Populations

	Data 1	Data 2	Data 3	Data 4
<i>N</i>	278	103	80	70
<i>n</i>	30	40	20	25
\bar{Y}	39.068	62.6212	51.8264	96.7000
\bar{X}	25.111	556.5541	11.2646	175.2671
<i>S_y</i>	56.4571 67	91.3549	18.3569	60.714
<i>S_x</i>	40.6747 97	610.1643	8.4563	140.8572
$\beta_{2(y)}$	25.8969	37.1279	2.2667	4.7596
$\beta_{2(x)}$	38.8898	17.8738	2.8664	7.0952
<i>h</i>	26.8142	17.2220	2.2209	4.6038
ρ	0.7213	0.7298	0.9413	0.7293

The biases and mean squared errors of the existing and proposed modified ratio type variance estimator for the populations given above are given in the following tables:

Table 2: Biases of the existing and proposed modified ratio type variance estimator

Estimator s	Bias			
	Data 1	Data 2	Data 3	Data 4
Isaki[1]	1282.9969 6	135.99347 1	10.87589 32	367.35093
Kadilar and Cingi[2]	1413.1211 6	136.16903 7	12.69804 37	367.80396 5
Kadilar and Cingi[2]	-138.3148	- 157.98237	33.32632 51	731.21715 1
Kadilar and Cingi[2]	1288.2034 5	136.00423 9	11.32844 5	367.40221 8
Kadilar and Cingi[2]	- 67.864364	- 178.34129	- 3.347314 2	-57.03954
Proposed Estimator	- 241.92096	- 784.70582	- 3.362946	- 77.110599

Table 3: MSE of the existing and proposed modified ratio type variance estimators

Estimator s	Mean Squared Error(MSE)			
	Data 1	Data 2	Data 3	Data 4
Isaki[1]	3778793.0 28	3579661 1.55	3924.948 201	1438805.63 7
Kadilar and Cingi[2]	3983111.9 77	3579672 0.6	4249.507 513	1439774.89 8
Kadilar and Cingi[2]	2514073.3 51	3582387 6.09	8667.305 927	2300502.29 5
Kadilar and Cingi[2]	3786820.8 36	3579661 8.23	4003.905 33	1438915.34 8
Kadilar and Cingi[2]	7990304.8 12	5983987 1.24	3636.702 702	1557091.88 4
Proposed Estimator	2614574.7 28	4533482 5.03	3750.092 953	1258334.94 1

From Table 2, it is observed that the bias of the proposed modified ratio type

Variance estimator is less than the biases of the traditional and existing modified ratio type variance estimators. Similarly from Table 3, it is observed that the mean squared error of the proposed modified ratio type variance estimator is less than the mean squared errors of the traditional and existing modified ratio type variance estimators.

6. CONCLUSION

In this article we have proposed a modified ratio type variance estimator using known value of population mean of auxiliary variable. The bias and mean squared error of the proposed modified ratio type variance estimator are obtained and compared with that of traditional ratio type variance estimator and existing modified ratio type variance estimators. We have also assessed the performances of the proposed estimator for some known populations. It is observed that the bias and mean squared error of the proposed estimator is less than the biases and mean squared errors of the traditional and existing estimators for certain known populations.

7. REFERENCES

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